

EnKF-based particle filters

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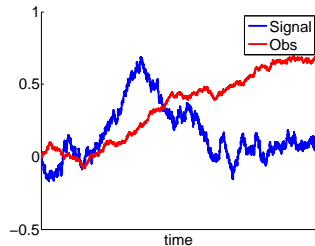
Filtering Problem

Signal

$$dX_t = f(X_t)dt + \sqrt{2}C dW_t$$

Observations

$$dY_t = h(X_t)dt + R^{1/2}dV_t.$$



Goal: determine

$$\pi(x|Y_{0:t})$$

State of the art

Linear Model: Kalman- Bucy Filter

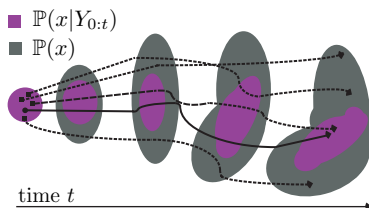
$$d\bar{x}_t^M = A\bar{x}_t^M dt + bdt - P_t^M H^T R^{-1}(H\bar{x}_t^M dt - dY_t)$$

Non-linear Model: approximate with empirical measure, i.e.,

$$\pi(x|Y_{0:t}) \approx \frac{1}{M} \sum_{i=1}^M \delta(x - X_t^i).$$

Ansatz: define modified evolution equation for particles X_t^i

Ensemble Kalman Filter (EnKF)



$$dX_t^i = f(X_t^i)dt + CdW_t^i - \frac{1}{2}P_t^M H^T R^{-1} (HX_t^i dt + H\bar{x}_t^M dt - 2dY_t)$$

$$\bar{x}_t^M = \frac{1}{M} \sum_{i=1}^M X_t^i \quad P_t^M = \frac{1}{M-1} \sum_{i=1}^M (X_t^i - \bar{x}_t^M)(X_t^i - \bar{x}_t^M)^T$$

Works remarkably well in practice: meteorology, oil reservoir exploration

But: theoretical understanding is largely missing

Recent accuracy results for EnKF

Interacting particle representation of the model error:

$$dX_t^i = f(X_t^i)dt + CC^\top (P_t^M)^{-1} (X_t^i - \bar{x}_t^M) dt \\ - \frac{1}{2} P_t^M H^\top R^{-1} (HX_t^i dt + H\bar{x}_t^M dt - 2dY_t)$$

Setting: $dY_t = X_t dt + R^{1/2} dW_t$ with $R = \varepsilon I$

Results: ([dWRS16])

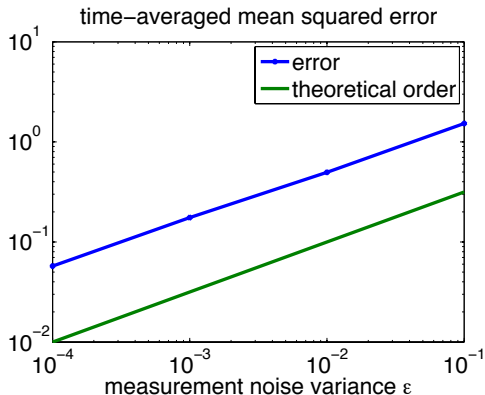
- ▶ Control of spectrum of covariance matrix P_t^M over $t \in [0, \infty)$.
- ▶ Control of estimation error $e_t = \|X_t^{\text{ref}} - \bar{x}_t^M\|^2$ in expectation

$$\mathbb{E}[E_t] = \mathcal{O}(\varepsilon^{1/2}) \quad (1)$$

and pathwise over fixed interval $t \in [0, T]$.

$$\mathbb{P}[\sup_{t \leq T} E_t \geq c_1 \varepsilon^q] \leq \mathcal{O}(\varepsilon^{1/2 - \eta - q}) \quad (2)$$

Numerical confirmation for L63



EnKF-based particle filters

Ansatz: modified evolution equation for the particles e.g., of the form

$$dX_t^i = f(X_t^i)dt + CdW_t^i - \sum_j X_t^j ds_t^{ji}$$

Aims:

- ▶ achieve first/second order accuracy
- ▶ to go beyond Gaussian assumption
- ▶ consistency for $M \rightarrow \infty$
- ▶ hybrid formulation to combine different interacting particle filters ([CRR16])

Linear ensemble transform filters (LETF)

Given: M samples $x_i^f \sim \pi(x)$ (prior)

Analysis step:

$$x_j^a = \sum_i x_i^f d_{ij} \quad (3)$$

with transformation matrix $D = \{d_{ij}\}$ subject to

$$\sum_{i=1}^M d_{ij} = 1 \quad (4)$$

Examples: ENKF, ESRF, NETF, ETPF (see [RC15])

Ensemble transform Particle Filter (ETPF)

Given:

- ▶ M samples $x_i^f \sim \pi(x)$ (prior ensemble)
- ▶ normalized importance weights $w_i \propto \pi(y|x_i^f)$ (likelihood)

Desired: M samples $x_i^a \sim \pi(x|y)$

Ansatz: replace resampling step with linear transformation by interpreting it as discrete Markov chain given by transition matrix

$$D \in \mathbb{R}^{M \times M}$$

s.t. $d_{ij} \geq 0$ and

$$\sum_i d_{ij} = 1, \quad \frac{1}{M} \sum_j d_{ij} = w_i.$$

Benefits: localization, hybrid

Ensemble transform Particle Filter (ETPF)

Then the posterior ensemble members are distributed according to the columns of the transformation

$$\tilde{X}_j^a \sim \begin{pmatrix} d_{1j} \\ d_{2j} \\ \vdots \\ d_{Mj} \end{pmatrix} \text{ and } x_j^a = \mathbb{E}[\tilde{X}_j^a] = \sum_i x_i^a d_{ij} \quad (5)$$

Example. Monomial resampling

$$D_{Mono} := \mathbf{w} \otimes \mathbf{1} = \begin{pmatrix} w_1 & w_1 & \cdots & w_1 \\ w_2 & w_2 & \cdots & w_2 \\ \vdots & \vdots & \ddots & \vdots \\ w_M & w_M & \cdots & w_M \end{pmatrix}.$$

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Here: X^f and X^a are independent.

Idea: increase correlation between X^f and X^a

Ensemble transform Particle Filter (ETPF)

Solve optimization problem

$$D_{ETPF} = \arg \min M \sum_{i,j=1}^M t_{ij} \|x_i^f - z_j^f\|^2$$

to find transformation matrix D_{ETPF} that increases correlation ([Rei13]).

Remarks:

- ▶ consistent for $M \rightarrow \infty$
- ▶ first-order accurate for finite M , i.e.,

$$\bar{x}^a = \frac{1}{M} \sum_{i=1}^M x_i^a = \sum_{i=1}^M w_i x_i^f \quad (6)$$

- ▶ **But:** not second-order accurate

Second-order accuracy

The analysis covariance matrix

$$\hat{P}^a = \frac{1}{M} \sum_{i=1}^M (x_i^a - \bar{x}^a)(x_i^a - \bar{x}^a)^T$$

is equal to the covariance matrix defined by the importance weights, i.e.

$$P^a = \sum_{i=1}^M w_i(t_k)(x_i^f - \bar{x}^a)(x_i^f - \bar{x}^a)^T.$$

First-order accurate LETFs

Notation:

$$\mathbf{X}^f = (x_1^f, \dots, x_M^f), \mathbf{X}^a = (x_1^a, \dots, x_M^a) \in \mathbb{R}^{N_x \times M}$$

and analogously

$$\mathbf{w} = (w_1, \dots, w_M)^T \in \mathbb{R}^{M \times 1} \text{ and } \mathbf{W} = \text{diag}(\mathbf{w})$$

LETF is first-order accurate if

$$\frac{1}{M} \mathbf{X}^a \mathbf{1} = \mathbf{X}^f \mathbf{w}.$$

This holds if D satisfies

$$\frac{1}{M} D \mathbf{1} = \mathbf{w}.$$

First-order accurate LETFs

Class of first-order accurate LETFs

$$\mathcal{D}_1 = \{D \in \mathbb{R}^{M \times M} \mid D^T \mathbf{1} = \mathbf{1}, D\mathbf{1} = M\mathbf{w}\}$$

Examples:

- ▶ $D_{\text{EnKF}} \notin \mathcal{D}_1$
- ▶ $D_{\text{ESRF}} \notin \mathcal{D}_1$
- ▶ $D_{\text{ETPF}} \in \mathcal{D}_1$
- ▶ $D_{\text{Mono}} \in \mathcal{D}_1$

Second-order accurate LETF

Analysis covariance matrix:

$$\hat{\mathbf{P}}^a = \frac{1}{M} \mathbf{X}^f (D - \mathbf{w}\mathbf{1}^T)(D - \mathbf{w}\mathbf{1}^T)^T (\mathbf{X}^f)^T \quad (7)$$

Importance sampling covariance matrix:

$$\mathbf{P}^a = \mathbf{X}^f (\mathbf{W} - \mathbf{w}\mathbf{w}^T) (\mathbf{X}^f)^T. \quad (8)$$

Thus the following equation has to hold for second-order accuracy:

$$(D - \mathbf{w}\mathbf{1}^T)(D - \mathbf{w}\mathbf{1}^T)^T = \mathbf{W} - \mathbf{w}\mathbf{w}^T \quad (9)$$

Class of second-order accurate LETFs

$$\mathcal{D}_2 = \{D \in \mathcal{D}_1 \mid (D - \mathbf{w}\mathbf{1}^T)(D - \mathbf{w}\mathbf{1}^T)^T = \mathbf{W} - \mathbf{w}\mathbf{w}^T\}. \quad (10)$$

Second-order corrected LETF

Given: $D \in \mathcal{D}_1$

Goal: correct transformation $\hat{D} \in \mathcal{D}_2$

Ansatz:

$$\hat{D} = D + \Delta$$

with $\Delta \in \mathbb{R}^{M \times M}$ such that $\Delta \mathbf{1} = \mathbf{0}$, $\Delta^T \mathbf{1} = \mathbf{0}$ ([dWAR17]).

Need to solve algebraic Riccati equation:

$$\begin{aligned} M(\mathbf{W} - \mathbf{w}\mathbf{w}^T) - (D - \mathbf{w}\mathbf{1}^T)(D - \mathbf{w}\mathbf{1}^T)^T \\ = (D - \mathbf{w}\mathbf{1}^T)\Delta^T + \Delta(D - \mathbf{w}\mathbf{1}^T)^T + \Delta\Delta^T. \end{aligned}$$

Numerical example I

Gaussian prior, non-Gaussian likelihood:

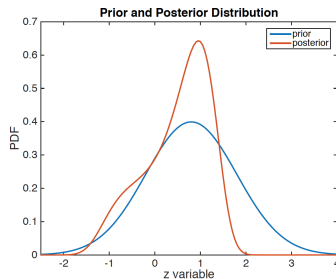


Figure: Prior and posterior distribution for the single Bayesian inference step

Numerical example I

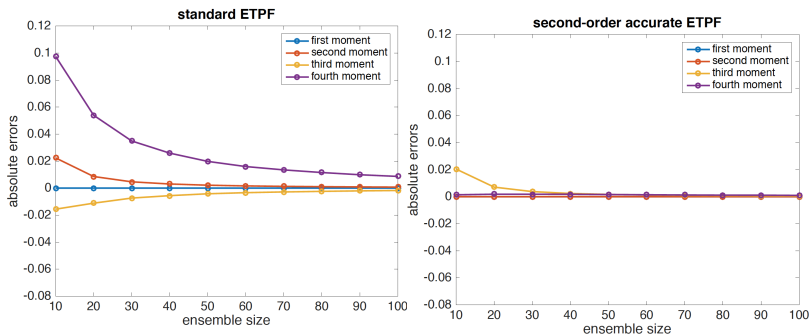


Figure: Absolute errors in the first four moments of the posterior distribution.

Numerical example II

Lorenz-63 model, first component observed infrequently ($\Delta t = 0.12$) and with large measurement noise ($R = 8$):

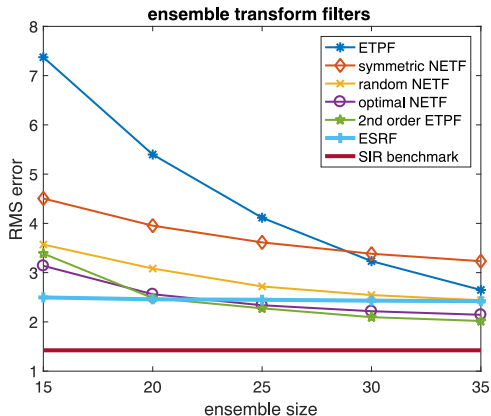


Figure: RMSEs for various second-order accurate LETFs compared to the ETPF, the ESRF, and the SIR PF as a function of the sample size, M .

Numerical example II

Hybrid filter: $D := D_{\text{ESRF}} D_{\text{ETPF}}$.

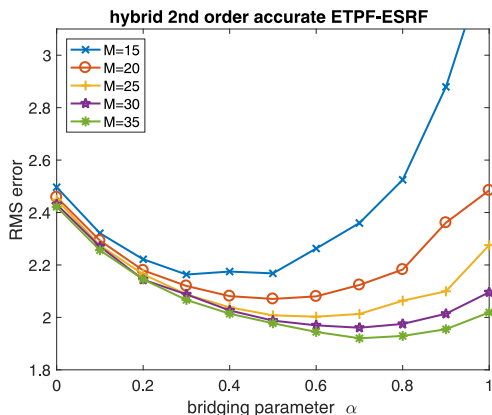


Figure: RMSEs for hybrid ESRF ($\alpha = 0$) and 2nd-order corrected LETF/ETPF ($\alpha = 1$) as a function of the sample size, M .



- ▶ PhD and Postdoc positions in 11 projects
- ▶ Mathematical foundation and applications in geophysics, biologie and cognitive sciences
- ▶ Support for external projects and long term visitors



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Thank you!

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