



# NON-LINEAR APPROXIMATION OF BAYESIAN UPDATE

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Figure : KAUST campus, 7 years old, approx. 7000 people (include 1700 kids), 100 nations, 36 km<sup>2</sup>.





## Advances in Uncertainty Quantification Methods, Algorithms and Applications (UQAW 2015)

January 6 – 9, 2015

9:00 a.m. – 5:00 p.m.

Level 0 auditorium, between Al-Jazri and Al-Kindi (buildings 4 and 5)

### WORKSHOP TOPICS

- 1- Uncertainty Quantification Methods and Algorithms
- 2- Verification and Validation
- 3- Experimental Design
- 4- Applications to Problems in Computational Science, Engineering, Networks and the Environment

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1. Computing the full Bayesian update is very expensive (MCMC is expensive)
2. Look for a cheap surrogate (linear, quadratic, cubic,... approx.)
3. Kalman filter is a particular case
4. Do Bayesian update of Polynomial Chaos Coefficients! (not probability densities!)
5. Consider non-Gaussian cases



1. O. Pajonk, B. V. Rosic, A. Litvinenko, and H. G. Matthies, A Deterministic Filter for Non-Gaussian Bayesian Estimation, Physica D: Nonlinear Phenomena, Vol. 241(7), pp. 775-788, 2012.
2. B. V. Rosic, A. Litvinenko, O. Pajonk and H. G. Matthies, Sampling Free Linear Bayesian Update of Polynomial Chaos Representations, J. of Comput. Physics, Vol. 231(17), 2012 , pp 5761-5787
3. A. Litvinenko and H. G. Matthies, Inverse problems and uncertainty quantification  
<http://arxiv.org/abs/1312.5048>, 2013
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General idea:

We **observe / measure** a system, whose structure we **know in principle**.

The system **behaviour** depends on some **quantities (parameters)**,  
which we **do not know**  $\Rightarrow$  **uncertainty**.

We model (uncertainty in) our **knowledge** in a **Bayesian** setting:  
as a **probability** distribution on the parameters.

We start with what we know **a priori**, then perform a **measurement**.

This gives new information, to **update** our knowledge  
(**identification**).

Update in probabilistic setting works with **conditional probabilities**  
 $\Rightarrow$  **Bayes's theorem**.

Repeated measurements lead to **better identification**.



Consider

$$A(u; q) = f \quad \Rightarrow \quad u = S(f; q),$$

where  $S$  is solution operator.

Operator depends on **parameters**  $q \in \mathcal{Q}$ ,  
hence state  $u \in \mathcal{U}$  is also function of  $q$ :

**Measurement** operator  $Y$  with values in  $\mathcal{Y}$ :

$$y = Y(q; u) = Y(q, S(f; q)).$$

Examples of measurements:

$$y(\omega) = \int_{\mathcal{D}_0} u(\omega, x) dx, \text{ or } u \text{ in few points}$$





For given  $f$ , measurement  $y$  is just a function of  $q$ .  
This function is usually **not invertible**  $\Rightarrow$  **ill-posed** problem,  
measurement  $y$  does **not** contain **enough information**.  
In **Bayesian** framework state of knowledge **modelled** in a  
probabilistic way,  
parameters  $q$  are **uncertain**, and **assumed** as **random**.  
**Bayesian** setting allows **updating / sharpening** of **information**  
about  $q$  when measurement is performed.  
The problem of updating **distribution**—**state of knowledge** of  $q$   
becomes **well-posed**.  
Can be applied **successively**, each new measurement  $y$  and  
forcing  $f$  —may also be uncertain—will provide **new**  
**information**.



With state  $\omega$  a RV, the quantity to be measured

$$y(\omega) = Y(q(\omega), u(\omega))$$

is also **uncertain**, a **random variable**.

Noisy data:  $\hat{y} + \epsilon(\omega)$ , where  $\hat{y}$  is the “true” value and a **random error**  $\epsilon$ .

Forecast of the measurement:  $z(\omega) = y(\omega) + \epsilon(\omega)$ .

Classically, **Bayes's theorem** gives **conditional probability**

$$\mathbb{P}(I_q | M_z) = \frac{\mathbb{P}(M_z | I_q)}{\mathbb{P}(M_z)} \mathbb{P}(I_q) \quad (\text{or } \pi_q(q | z) = \frac{p(z | q)}{Z_s} p_q(q))$$

expectation with this posterior measure is **conditional expectation**. **Kolmogorov** starts from **conditional expectation**  
 $\mathbb{E}(\cdot | M_z)$ ,

from this **conditional probability** via  $\mathbb{P}(I_q | M_z) = \mathbb{E}(\chi_{I_q} | M_z)$ .



The **conditional expectation** is **defined** as **orthogonal projection** onto the closed **subspace**  $L_2(\Omega, \mathbb{P}, \sigma(z))$ :

$$\mathbb{E}(q|\sigma(z)) := P_{\mathcal{L}_\infty} q = \operatorname{argmin}_{\tilde{q} \in L_2(\Omega, \mathbb{P}, \sigma(z))} \|q - \tilde{q}\|_{L_2}^2$$

The subspace  $\mathcal{L}_\infty := L_2(\Omega, \mathbb{P}, \sigma(z))$  represents the **available information**.

The **update**, also called the **assimilated** value  $q_a(\omega) := P_{\mathcal{L}_\infty} q = \mathbb{E}(q|\sigma(z))$ , and represents **new state** of knowledge **after** the measurement.

Doob-Dynkin:  $\mathcal{L}_\infty = \{\varphi \in \mathcal{L} : \varphi = \phi \circ z, \phi \text{ measurable}\}$ .



Multivariate Hermite polynomials were used to approximate random fields/stochastic processes with Gaussian random variables. According to Cameron and Martin theorem PCE expansion converges in the  $L_2$  sense.

Let  $Y(x, \theta)$ ,  $\theta = (\theta_1, \dots, \theta_M, \dots)$ , is approximated:

$$Y(x, \theta) = \sum_{\beta \in \mathcal{J}_{m,p}} H_{\beta}(\theta) Y_{\beta}(x), \quad |\mathcal{J}_{m,p}| = \frac{(m+p)!}{m!p!},$$

$$H_{\beta}(\theta) = \prod_{k=1}^M h_{\beta_k}(\theta_k),$$

$$Y_{\beta}(x) = \frac{1}{\beta!} \int_{\Theta} H_{\beta}(\theta) Y(x, \theta) \mathbb{P}(d\theta).$$

$$Y_{\beta}(x) \approx \frac{1}{\beta!} \sum_{i=1}^{N_q} H_{\beta}(\theta_i) Y(x, \theta_i) w_i.$$



Look for  $\varphi$  such that  $q(\xi) = \varphi(z(\xi))$ ,  $z(\xi) = y(\xi) + \varepsilon(\omega)$ :

$$\varphi \approx \tilde{\varphi} = \sum_{\alpha \in \mathcal{J}_p} \varphi_\alpha \Phi_\alpha(z(\xi)) \quad (1)$$

and minimize  $\|q(\xi) - \tilde{\varphi}(z(\xi))\|^2$ , where  $\Phi_\alpha$  are polynomials (e.g. Hermite, Laguerre, Chebyshev or something else). Taking derivatives with respect to  $\varphi_\alpha$ :

$$\frac{\partial}{\partial \varphi_\alpha} \langle q(\xi) - \tilde{\varphi}(z(\xi)), q(\xi) - \tilde{\varphi}(z(\xi)) \rangle = 0 \quad \forall \alpha \in \mathcal{J}_p \quad (2)$$

Inserting representation for  $\tilde{\varphi}$ , obtain



$$\begin{aligned}
 & \frac{\partial}{\partial \varphi_\alpha} \mathbb{E} \left( q^2(\xi) - 2 \sum_{\beta \in \mathcal{J}} q \varphi_\beta \Phi_\beta(z) + \sum_{\beta, \gamma \in \mathcal{J}} \varphi_\beta \varphi_\gamma \Phi_\beta(z) \Phi_\gamma(z) \right) \\
 &= 2 \mathbb{E} \left( -q \Phi_\alpha(z) + \sum_{\beta \in \mathcal{J}} \varphi_\beta \Phi_\beta(z) \Phi_\alpha(z) \right) \\
 &= 2 \left( \sum_{\beta \in \mathcal{J}} \mathbb{E} [\Phi_\beta(z) \Phi_\alpha(z)] \varphi_\beta - \mathbb{E} [q \Phi_\alpha(z)] \right) = 0 \quad \forall \alpha \in \mathcal{J}
 \end{aligned}$$

$$\mathbb{E} [\Phi_\beta(z) \Phi_\alpha(z)] \varphi_\beta = \mathbb{E} [q \Phi_\alpha(z)]$$



Now, rewriting the last sum in a matrix form, obtain the linear system of equations ( $=: \mathbf{A}$ ) to compute coefficients  $\varphi_\beta$ :

$$\begin{pmatrix} \dots & \dots & \dots \\ \vdots & \mathbb{E}[\Phi_\alpha(z(\xi))\Phi_\beta(z(\xi))] & \vdots \\ \dots & \dots & \dots \end{pmatrix} \begin{pmatrix} \vdots \\ \varphi_\beta \\ \vdots \end{pmatrix} = \begin{pmatrix} \vdots \\ \mathbb{E}[q(\xi)\Phi_\alpha(z(\xi))] \\ \vdots \end{pmatrix},$$

where  $\alpha, \beta \in \mathcal{J}$ ,  $\mathbf{A}$  is of size  $|\mathcal{J}| \times |\mathcal{J}|$ .



Using the same quadrature rule of order  $q$  for each element of  $A$ , we can write

$$A = \mathbb{E} \left[ \Phi_{\mathcal{J}_\alpha}(z(\xi)) \Phi_{\mathcal{J}_\beta}(z(\xi))^T \right] \approx \sum_{i=1}^{N^A} w_i^A \Phi_{\mathcal{J}_\alpha}(z_i) \Phi_{\mathcal{J}_\beta}(z_i)^T, \quad (3)$$

where  $(w_i^A, \xi_i)$  are weights and quadrature points,  $z_i := z(\xi_i)$  and  $\Phi_{\mathcal{J}_\alpha}(z_i) := (\dots \Phi_\alpha(z(\xi_i)) \dots)^T$  is a vector of length  $|\mathcal{J}_\alpha|$ .

$$b = \mathbb{E} [q(\xi) \Phi_{\mathcal{J}_\alpha}(z(\xi))] \approx \sum_{i=1}^{N^b} w_i^b q(\xi_i) \Phi_{\mathcal{J}_\alpha}(z_i), \quad (4)$$

where  $\Phi_{\mathcal{J}_\alpha}(z(\xi_i)) := (\dots, \Phi_\alpha(z(\xi_i)), \dots)$ ,  $\alpha \in \mathcal{J}_\alpha$ .





We can write the Eq. 15 with the right-hand side in Eq. 4 in the compact form:

$$[\Phi_A] [\text{diag}(\dots w_i^A \dots)] [\Phi_A]^T \begin{pmatrix} \vdots \\ \varphi_\beta \\ \vdots \end{pmatrix} = [\Phi_b] \begin{pmatrix} w_0^b q(\xi_0) \\ \dots \\ w_{N^b}^b q(\xi_{N^b}) \end{pmatrix} \quad (5)$$

$$[\Phi_A] \in \mathbb{R}^{\mathcal{J}_\alpha \times N^A}, [\text{diag}(\dots w_i^A \dots)] \in \mathbb{R}^{N^A \times N^A}, [\Phi_b] \in \mathbb{R}^{\mathcal{J}_\alpha \times N^b},$$

$$[w_0^b q(\xi_0) \dots w_{N^b}^b q(\xi_{N^b})] \in \mathbb{R}^{N^b}.$$

Solving Eq. 5, obtain vector of coefficients  $(\dots \varphi_\beta \dots)^T$  for all  $\beta$ .  
Finally, the assimilated parameter  $q_a$  will be

$$q_a = q_f + \tilde{\varphi}(\hat{y}) - \tilde{\varphi}(z), \quad (6)$$

$$z(\xi) = y(\xi) + \varepsilon(\omega), \quad \tilde{\varphi} = \sum_{\beta \in \mathcal{J}_p} \varphi_\beta \Phi_\beta(z(\xi))$$



Assume  $z(\xi) = \xi^2$  and  $q(\xi) = \xi^3$ . The normalized PCE coefficients are  $(1, 0, 1, 0)$

$$(\xi^2 = 1 \cdot H_0(\xi) + 0 \cdot H_1(\xi) + 1 \cdot H_2(\xi) + 0 \cdot H_3(\xi))$$

and  $(0, 3, 0, 1)$

$$(\xi^3 = 0 \cdot H_0(\xi) + 3 \cdot H_1(\xi) + 0 \cdot H_2(\xi) + 1 \cdot H_3(\xi)).$$

For such data the **mapping  $\varphi$  does not exist**. The matrix  $A$  is close to singular.

Support of Hermite polynomials (used for Gaussian RVs) is  $(-\infty, \infty)$ .



Assume  $z(\xi) = \xi^2$  and  $q(\xi) = \xi^3$ .

The normalized gPCE coefficients are  $(2, -4, 2, 0)$  and  $(6, -18, 18, -6)$ .

For such data the **mapping mapping  $\varphi$  of order 8 and higher produces a very accurate result.**

Support of Laguerre polynomials (used for Gamma RVs) is  $[0, \infty)$ .



Is a system of ODEs. Has chaotic solutions for certain parameter values and initial conditions.

$$\dot{x} = \sigma(\omega)(y - x)$$

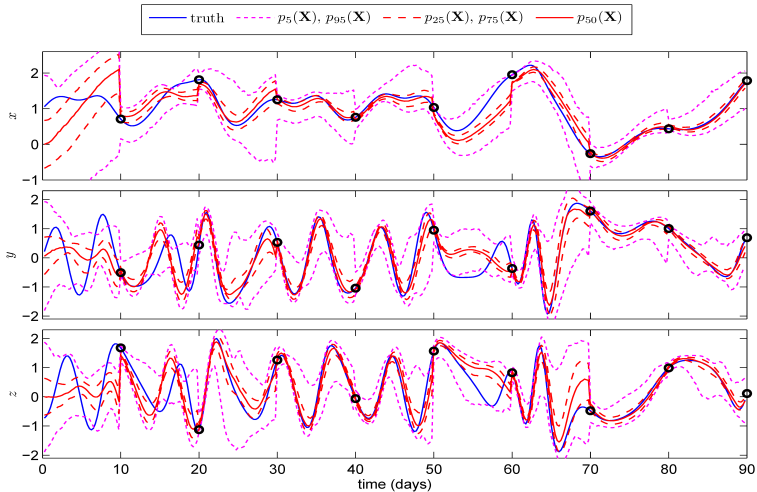
$$\dot{y} = x(\rho(\omega) - z) - y$$

$$\dot{z} = xy - \beta(\omega)z$$

Initial state  $q_0(\omega) = (x_0(\omega), y_0(\omega), z_0(\omega))$  are uncertain.

Solving in  $t_0, t_1, \dots, t_{10}$ , **Noisy Measur. → UPDATE**, solving in  $t_{11}, t_{12}, \dots, t_{20}$ , **Noisy Measur. → UPDATE**,...





Trajectories of  $x, y$  and  $z$  in time. After each update (new information coming) the uncertainty drops. (O. Pajonk)

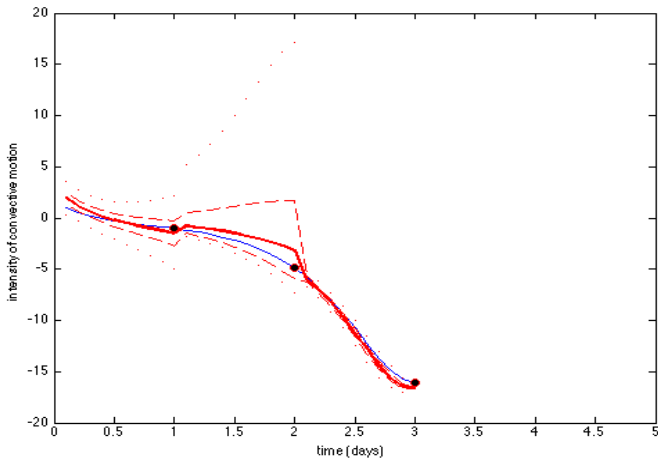


Figure : Partial state trajectory with uncertainty and three updates

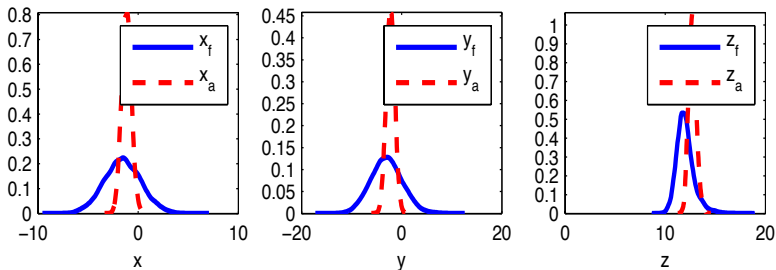


Figure : NLBU: Linear measurement  $(x(t), y(t), z(t))$ : prior and posterior after one update

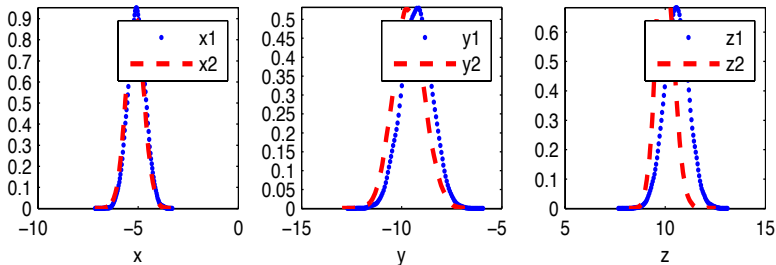


Figure : Linear measurement: Comparison posterior for LBU and NLBU after second update



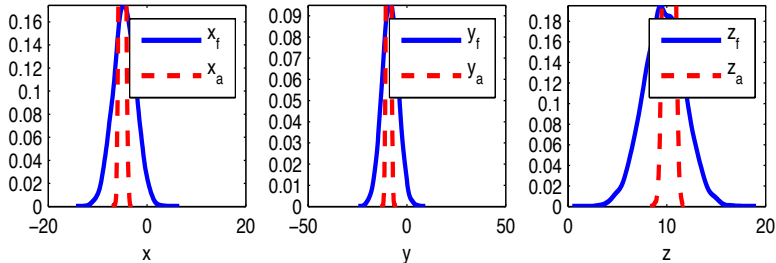


Figure : Quadratic measurement ( $x(t)^2, y(t)^2, z(t)^2$ ): Comparison of a priori and a posteriori for NLBU



Taken from Stochastic Galerkin Library (sglib), by Elmar Zander (TU Braunschweig)

$$-\nabla \cdot (\kappa(x, \xi) \nabla u(x, \xi)) = f(x, \xi), \quad x \in [0, 1]$$

Measurements are taken at  $x_1 = 0.2$ , and  $x_2 = 0.8$ . The means are  $\bar{y}(x_1) = 10$ ,  $\bar{y}(x_2) = 5$  and the variances are 0.5 and 1.5 correspondingly.

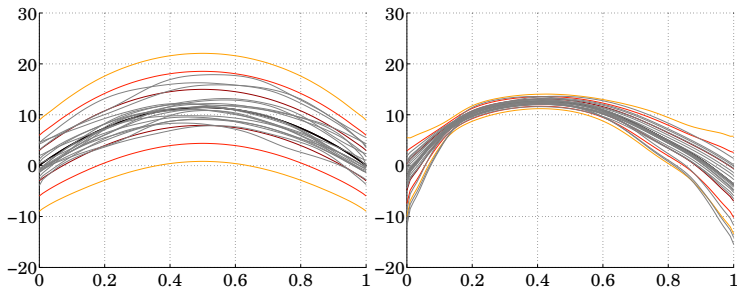


Figure : Original and updated solutions, mean value plus/minus 1,2,3 standard deviations

See more in [sglib](#) by [Elmar Zander](#)

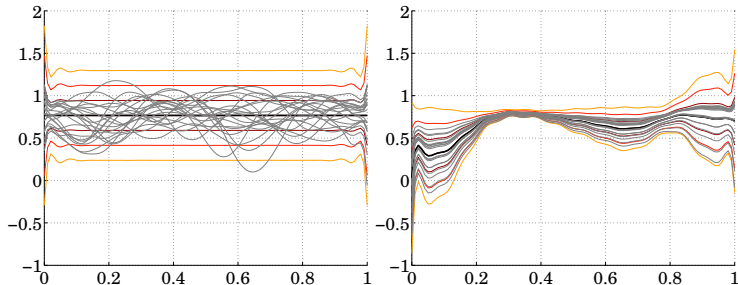


Figure : Original and updated parameter  $q$ .

See more in [sglib](#) by [Elmar Zander](#)



- ▶ + Step 1. Introduced a way to derive MMSE  $\varphi$  (as a linear, quadratic, cubic etc approximation, i. e. compute conditional expectation of  $q$ , given measurement  $Y$ ).
- ▶ Step 2. Apply  $\varphi$  to identify parameter  $q$
- ▶ + All ingredients can be given as gPC.
- ▶ + we apply it to solve inverse problems (ODEs and PDEs).
- ▶ - Stochastic dimension grows up very fast.



I used a Matlab toolbox for stochastic Galerkin methods (sglib)

<https://github.com/ezander/sglib>

Alexander Litvinenko and his research work was supported by the King Abdullah University of Science and Technology (KAUST), SRI-UQ and ECRC centers.



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