

Uncertainty quantification for source reconstruction of ^{137}Cs released during the Fukushima accident

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EnKF workshop 2022 - Norce

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Lundi 30 mai, 2022

Fukushima-Daiichi accident

Fukushima-Daiichi nuclear disaster in March 2011.

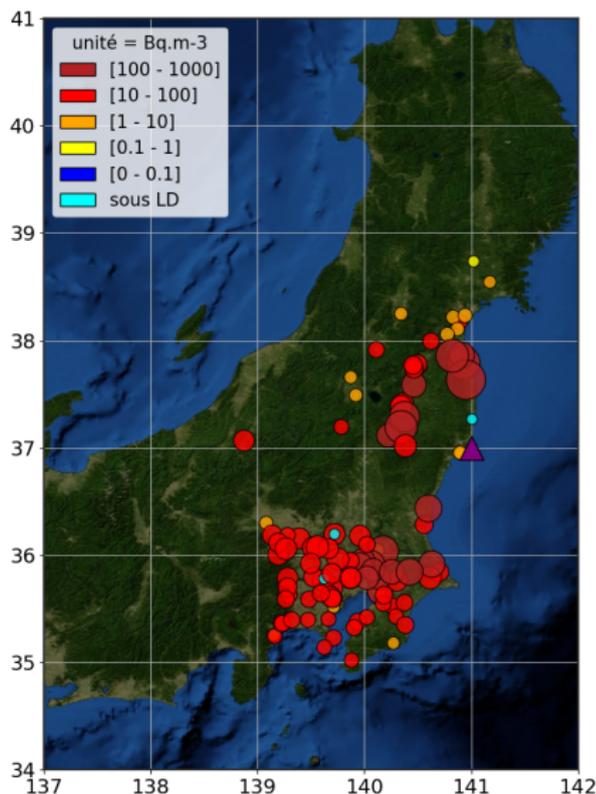


Release of large quantities of radionuclides, including ^{137}Cs :

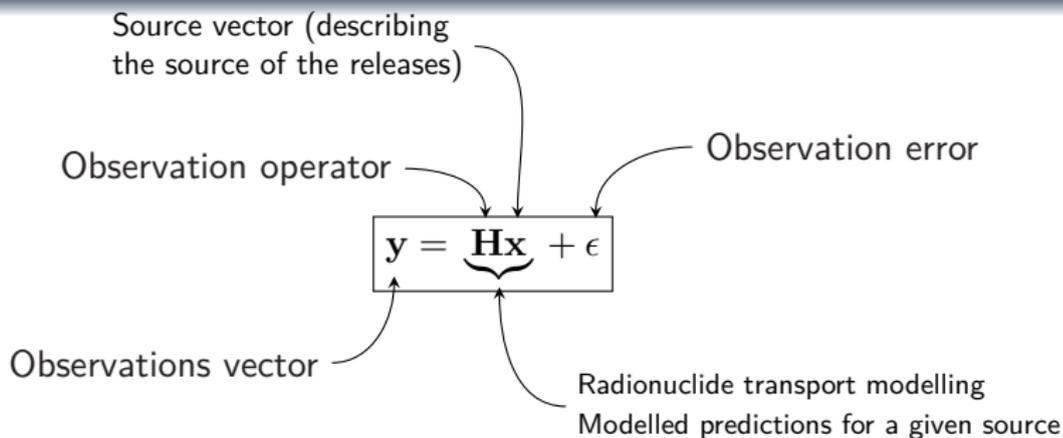
- on three weeks,
- with an important temporal variability.

Fukushima-Daiichi ^{137}Cs observations

- > 14,000 hourly air concentration measurements between 11/03/2011 and 23/03/2011
- > 1,000 deposition measurements;
- Use of the Eulerian transport model IdX represented by a linear observation operator \mathbf{H}
- Meteorological data: ECMWF OD (0.125° , 3h);



Observation equation

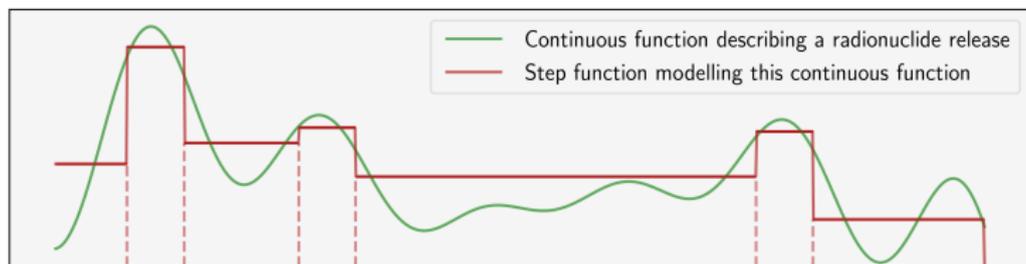


- All variables describing the source of the release (height, coordinates,...) are assumed to be known, except for
 - ▶ the **source term** \mathbf{q} or vector of constant release rates of size N_{imp} (number of pulses).

But how to characterise \mathbf{q} ?

Representation of the continuous release by a step function

The source term q needs to be well characterised



Representation of the release \rightarrow solving a trade-off between

- bias (too simple model, insufficient to learn correctly from data),
- variance (overfitting or overinterpretation of the data).

Inverse the source \rightarrow inverse the parametrisation of the source

Bayesian inverse modelling

Bayes' formula

Bayes' formula, with \mathbf{x} the vector of variables characterising the source and \mathbf{y} the observations is written

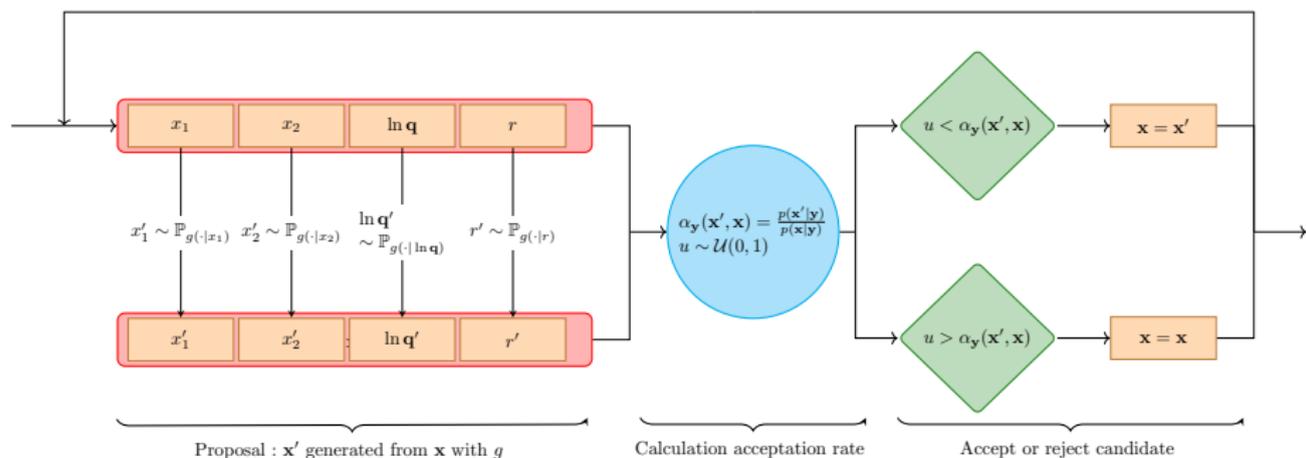
$$\underbrace{p(\mathbf{x}|\mathbf{y})}_{\text{Posterior}} = \frac{\overbrace{p(\mathbf{y}|\mathbf{x})}^{\text{Likelihood}} \overbrace{p(\mathbf{x})}^{\text{Prior}}}{\underbrace{p(\mathbf{y})}_{\text{Evidence}}} \propto p(\mathbf{y}|\mathbf{x})p(\mathbf{x}).$$

$\mathbf{y}|\mathbf{x}$ diagnostics the difference between the observations \mathbf{y} and the dispersion model results computed out of the source \mathbf{x} .

Source vector = variables of interest to sample:

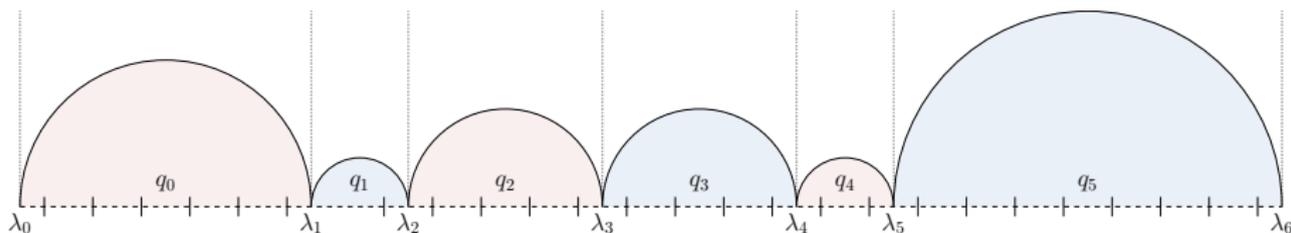
- release rates \mathbf{q} ;
- observation error scale matrix \mathbf{R} ;
- source prior scale matrix \mathbf{B} .

A popular MCMC algorithm: Metropolis-Hastings



Model selection: Reversible-Jump MCMC

- The constant release rates q_i are separated by "edges" λ_i .
- The evolution of the release is modelled by a specific *partition* of edges.



The transdimensional partition of edges $\Lambda = \{\lambda_0, \lambda_1, \dots\}$ is integrated as a variables ensemble to the MCMC procedure:

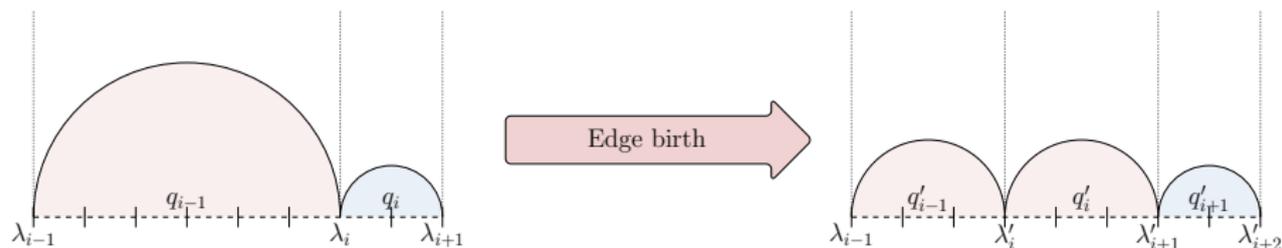
$$\mathbf{x} = (\ln \mathbf{q}, \mathbf{R}) \rightarrow \mathbf{x} = (\ln \mathbf{q}, \mathbf{R}, \Lambda)$$

\Rightarrow Use of the Reversible Jump MCMC¹.

¹Green 1995; Liu et al. 2017.

RJ-MCMC: random transdimensional jumps

An example of transdimensional procedure : the "edge" birth².



$$\mathbf{x} = (\lambda_0, \dots, \lambda_n, q_0, \dots, q_k, \dots) \longleftrightarrow (\lambda_0, \dots, \lambda'_i, \dots, \lambda_n, q_0, \dots, q_k - u_{\ln q}, q_k + u_{\ln q}, \dots) = \mathbf{x}'$$

- Dimension change of $\mathbf{x} \rightarrow$ addition of terms ensuring the *detailed balance* in the MCMC algorithm;
- Need to define new priors and transition probabilities on the edge variables and release rates.

²Bodin and Sambridge 2009.

Redefining the observation error scale matrix

- Need to integrate more information.
 - ▶ Factor both concentration and deposition measurements into a Bayesian sampling
 - ▶ Take in account spatial distances between air concentration measurements

$$\mathbf{R} = \begin{bmatrix} r_{c,i} & 0 & \dots & \dots & \dots & 0 \\ 0 & r_{c,j} & 0 & \dots & \dots & \dots \\ \dots & 0 & r_{c,i} & \dots & \dots & \dots \\ \dots & \dots & 0 & \dots & 0 & \dots \\ \dots & \dots & \dots & \dots & r_d & 0 \\ 0 & \dots & \dots & \dots & 0 & r_d \end{bmatrix}. \quad (1)$$

Vector to sample:

$$\mathbf{x} = (\ln \mathbf{q}, r_{c,1}, \dots, r_{c,i}, \dots, r_d, (\lambda_1, \dots, \lambda_{N_{\text{imp}}-1})) \quad (2)$$

Definition of the distributions

- **Likelihood** $\mathbf{y}|\mathbf{x} \sim \text{log-Cauchy}$ with scale \mathbf{R}^3 .

$$p(\mathbf{y}|\mathbf{x}) \propto \prod_{i=1}^{\text{Nobs}} \frac{1}{(y_i + y_t)\pi\sqrt{r_i} \left(1 + \frac{(\ln(y_i + y_t) - \ln(\mathbf{H}\mathbf{x}_i + y_t))^2}{r_i}\right)} \quad (3)$$

- ▶ efficient to manage observations of different orders of magnitude;
- ▶ to manage the observations close to zero \rightarrow we add a threshold term⁴ y_t ;

- **Prior** definitions

- ▶ Uniform priors on the scale parameters;
- ▶ Exponential prior on edges to penalise too complex models:

$$p(\lambda_1, \dots, \lambda_k) = \begin{cases} \frac{e^{-k}}{\sum_{i=1}^{N_{b,\max}} \frac{N_{b,\max}!}{i!(N_{b,\max}-i)!} e^{-i}}, & \text{if } k \in \{1, 2, \dots, N_{b,\max}\}; \\ 0 & \text{otherwise,} \end{cases} \quad (4)$$

³Dumont Le Brazidec et al. 2021.

⁴Liu et al. 2017.

Prior on the release rates

- Some release pulses are not constrained by the observations.
- Folded gaussian prior is set with \mathbf{B} parametrised with parameters b_c for constrained pulses and b_{nc} for non-constrained pulses
- \mathbf{B} is adapted for our case: pulses sampled are combinations of hourly pulses

$$p(\ln \mathbf{q} | N_{\text{imp}}) = \prod_{i=1}^{N_{\text{imp}}} \sqrt{\frac{2}{\pi(w_{c,i}b_c + w_{nc,i}b_{nc})}} \left(e^{-\frac{(\ln q_i)^2}{2(w_{c,i}b_c + w_{nc,i}b_{nc})}} \right)$$

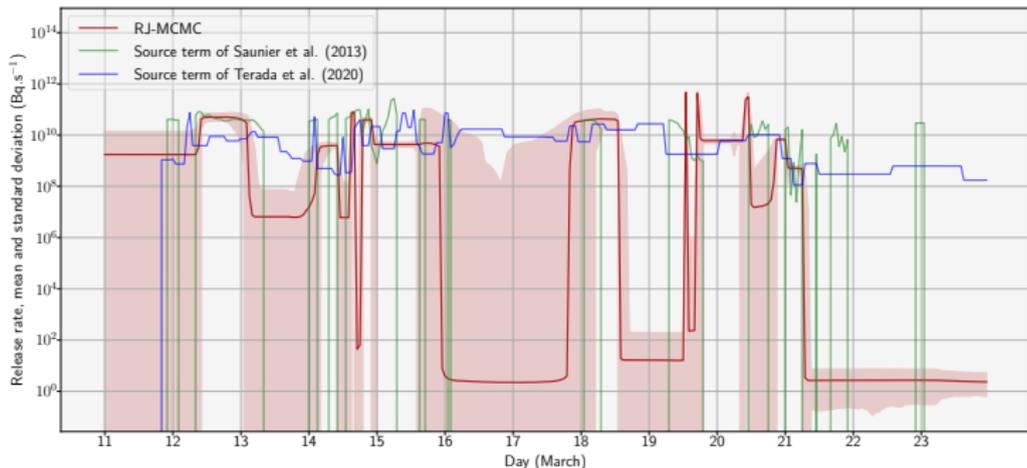
- N_{imp} is the number of pulses at a certain RJ-MCMC iteration
 - ▶ characterises the grid on which $\ln \mathbf{q}$ is defined.
 - ▶ $N_{\text{imp}} = N_b - 1$ (N_b is the number of edges)
- this prior also constrains the model's complexity

We sample:

$$\mathbf{x} = (\ln \mathbf{q}, r_{c,1}, \dots, r_{c,i}, \dots, r_d, b_c, (\lambda_1, \dots, \lambda_{N_{\text{imp}}-1})) \quad (5)$$

Fukushima-Daiichi ^{137}Cs release rate reconstruction

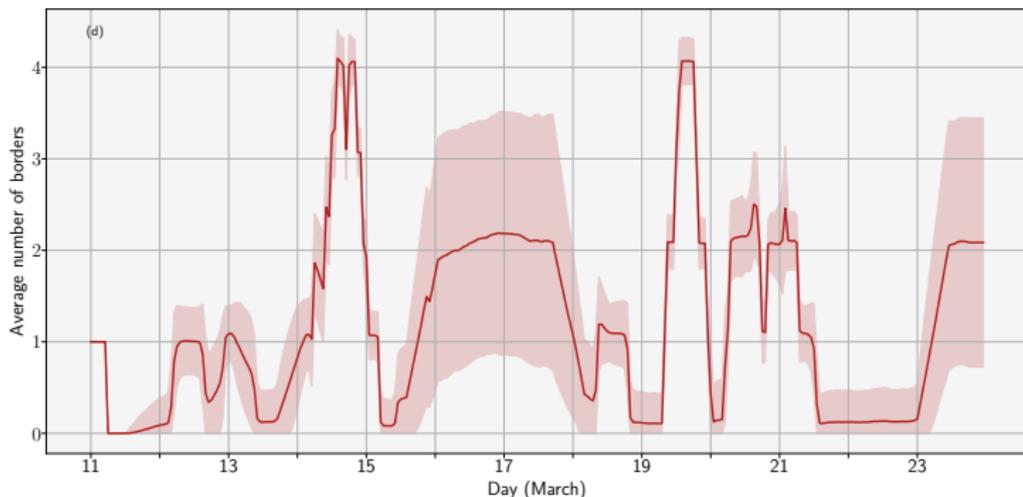
Fukushima-Daiichi ^{137}Cs release rate evolution (and corresponding variance) in $\text{Bq}\cdot\text{s}^{-1}$.



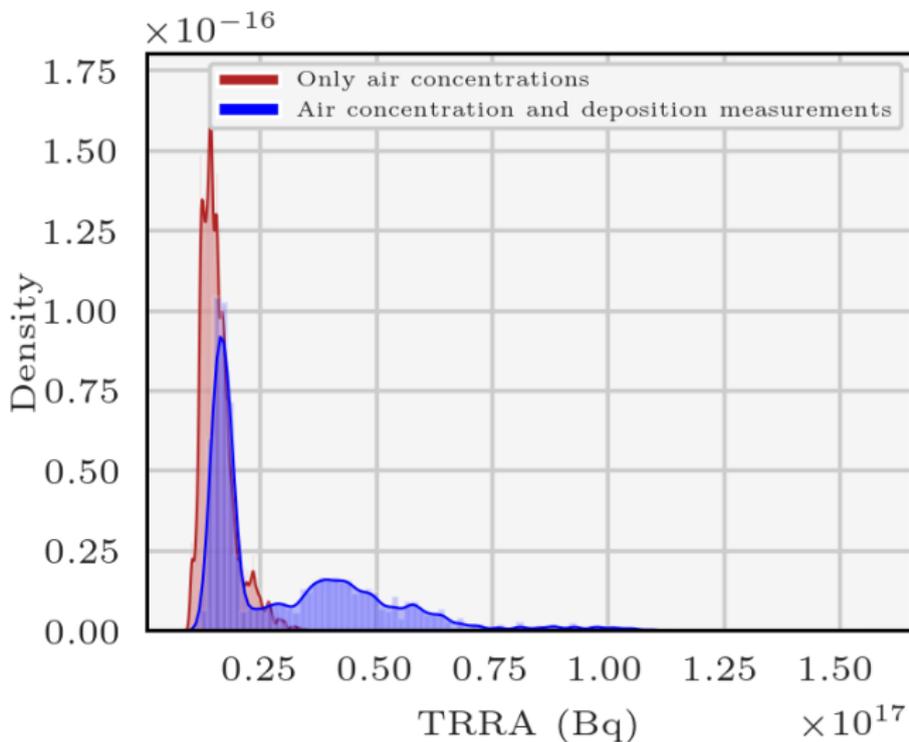
- Several release episodes (12-13 march, 14-16 march, 18 march, 19-21 march) of diverse variabilities;
- Large *variability of the variability* → proves RJ-MCMC pertinence.

Fukushima-Daiichi ^{137}Cs variability reconstruction

Evolution of the number of edges sampled around each hour (and corresponding variance).



- High variability between march 14 and 15, and march 19 and 21: periods of intense release;
- Low variability elsewhere (apart from artefacts in non-constraints periods).

Fukushima-Daiichi ^{137}Cs source term reconstructionReconstruction of the total ^{137}Cs release with or without deposition measurements

Conclusions

- Complex releases \rightarrow high variability and high variability of the variability
- Model of such a release complicated to define because:
 - ▶ might be at some periods too simple
 \rightarrow *bias* errors
 - ▶ might be at some periods too complicated
 \rightarrow overfitting + *variance* errors;
- Use of RJ-MCMC allows to
 - ▶ reconstruct the best model by solving the bias-variance trade-off,
 - ▶ and thus, better estimate the uncertainties related to the release representation.