

# Can a training image be a substitute for a random field model?

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# Introduction

Modern stochastic data assimilation algorithms may require generating ensembles of facies fields. This is typically the case in reservoir optimization where each facies field is used as input for a fluid flow exercise.

In a geostatistical context, facies fields are nothing but conditional simulations. Different approaches can be considered to produce them:

- By resorting to a **spatial stochastic model** such as the plurigaussian model, the Boolean model... This requires the choice of a model, the statistical inference of its parameters, the design of a conditional simulation algorithm...
- By resorting to a **training image** to produce multipoint simulations (MPS): no statistical inference, wide generality, conceptual simplicity...

The second approach looks miraculous. Isn't there a price to pay for it?

# Outline

## Compatibility between MPS's and stochastic simulations

- Principle of MPS
- Case of an infinite training image
- Case of a finite training image

## Statistical considerations on template matching

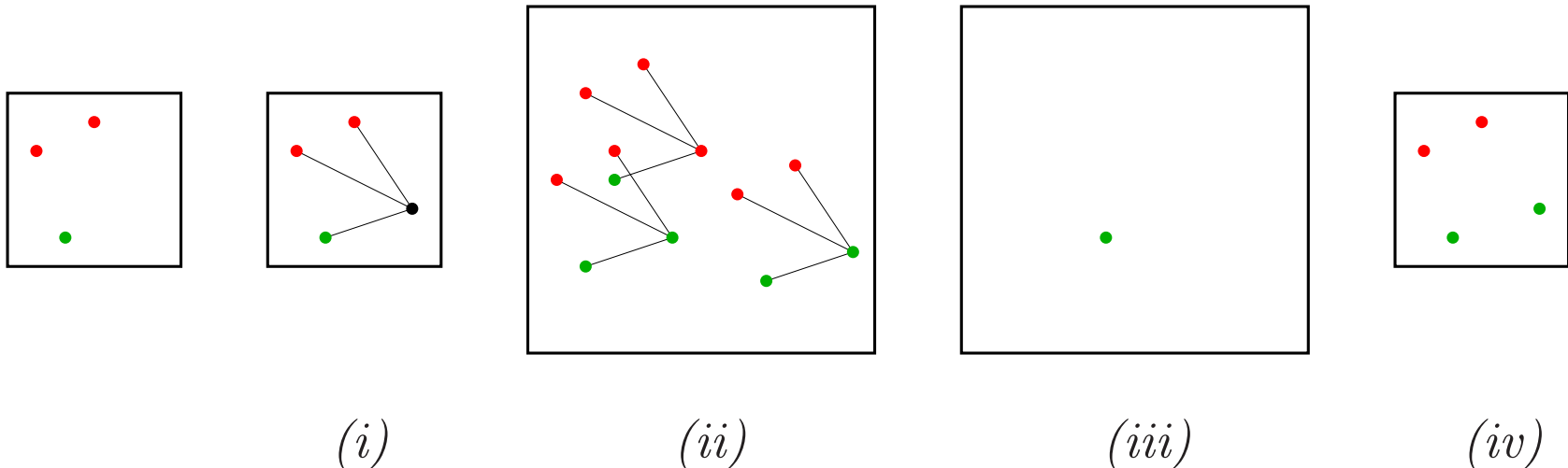
- Statistical matching of a template
- Application to the estimation of the size of a training image
- Example
- A simple combinatorial remark

# **Compatibility between MPS's and stochastic simulations**

# Principle of MPS

This is a **sequential algorithm**. Each step is as follows:

- (i) a new target point is selected at random in the simulation field. It defines a **template** along with the already processed points;*
- (ii) the pixels where the template **matches** the training image are identified;*
- (iii) one pixel among those is selected at random;*
- (iv) its value is assigned to the target point.*



# The problem addressed

## Assumption:

Suppose that the training image  $I$  is a realization, or part of a realization, of some **stationary, ergodic random field** (SERF)  $Z$  on  $\mathbb{Z}^2$ .

$Z$  is ergodic means that its spatial distribution can be retrieved from any of its realizations:

$$P\{\cap_{i=1,n} Z(x_i) = \epsilon_i\} = \lim_{S \rightarrow \mathbb{Z}^2} \frac{1}{\#S} \sum_{s \in S} \prod_{i=1}^n 1_{I(x_i+s)=\epsilon_i}$$

## Question:

Does the empirical spatial distribution yielded by MPS's fit that of  $Z$ ?

# Case of an infinite training image

## Remark:

The algorithm cannot be directly applied because the template  $T$  matches  $I$  at infinitely many points (set  $S_T$ ). The target point is then assigned the value 0 or 1 with respective probabilities

$$p_0 = \lim_{S \rightarrow \mathbb{Z}^2} \frac{1}{\#S} \sum_{s \in S \cap S_T} 1_{I(s)=0} \quad p_1 = \lim_{S \rightarrow \mathbb{Z}^2} \frac{1}{\#S} \sum_{s \in S \cap S_T} 1_{I(s)=1}$$

## Results:

– Each MPS is a **patch** of the TI;

– The empirical **spatial distribution** fits that of  $Z$ :

If  $(X_k, k \geq 1)$  is a sequence of MPS's on domain  $D$ , if  $x_1, \dots, x_n \in D$  and if  $\epsilon_1, \dots, \epsilon_n \in \{0, 1\}$ , then

$$= \lim_{k \rightarrow \infty} \frac{1}{k} \sum_{\ell=1}^k \prod_{i=1}^n 1_{X_\ell(x_i)=\epsilon_i} = P\{\cap_{i=1,n} Z(x_i) = \epsilon_i\}$$

– **Conditional MPS** can be performed as well.

# Case of a finite training image

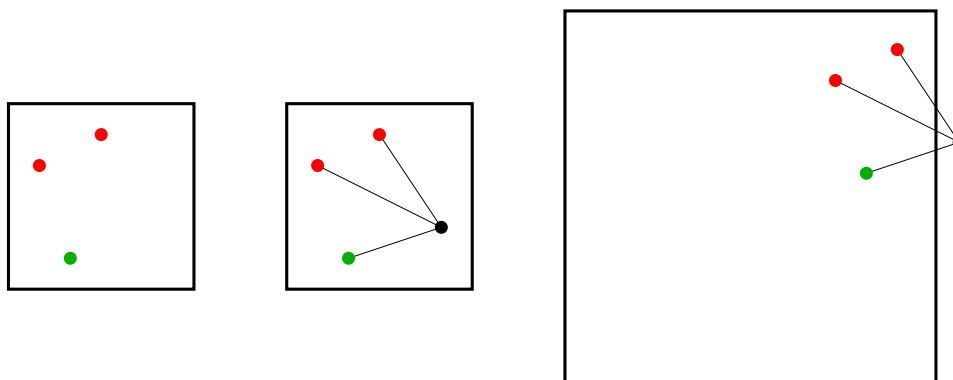
## Uncommon situation:

The algorithm runs till a MPS has been completed:

- Then the MPS a **patch** of the training image;
- Different MPS's display little variability (the training image has less variability than an entire realization, possible overlaps between MPS's).

## Common situation:

The algorithm stops at one step because the training image does not match the template at any location:





# How to prevent the algorithm from stopping?

## Reduce the size of the template

- By discarding points of a template, spurious **conditional independence relationships** are introduced (Holden, 2006);
- Because of the sequential nature of the algorithm, these relationships propagate, which may lead to **severe artefacts** to the final outcome (Arpat, 2005).

## Increase the size of the training image

- MPS algorithms works for infinitely large images
- Accordingly, it should also work provided that the training image is **large enough...**

# **Statistical considerations on template matching**

# Statistical matching of a template

## Notation:

- $Z$  is a binary, stationary, ergodic random field (SERF) on  $\mathbb{Z}^2$ ;
- $T$  is a template.

## Matching:

Let  $N_T(x) = 1$  if the template located at  $x$  matches  $Z$ , and 0 otherwise.  $N_T$  is also a SERF. Its mean, variance and correlation function are respectively denoted by  $\mu_T$ ,  $\sigma_T^2 = \mu_T(1 - \mu_T)$  and  $\rho_T$ .

## Matching number:

More generally, the number of times  $T$  matches  $Z$  in a finite domain  $V$  is  $N_T(V) = \sum_{x \in V} N_T(x)$ . We have ( $\tau_h$  is the translation by vector  $o\vec{h}$ )

$$E\{N_T(V)\} = \mu_T \#V$$

$$Var\{N_T(V)\} = \sigma_T^2 \sum_{h \in \mathbb{Z}^2} \rho_T(h) \#(V \cap \tau_h V)$$

# An asymptotic result

Heuristic approach:

$$\text{Var}\{N_T(V)\} = \sigma_T^2 \sum_{h \in \mathbb{Z}^2} \rho_T(h) \#(V \cap \tau_h V)$$

If the range of  $\rho_T$  is small compared to the size of  $V$ , then one heuristically has  $\#(V \cap \tau_h V) \approx \#V$  whenever  $\rho_T \not\approx 0$ , which implies

$$\text{Var}\{N_T(V)\} \approx \sigma_T^2 \sum_{h \in \mathbb{Z}^2} \rho_T(h) \#V$$

**Definition:**

The integral  $a_T = \sum_{h \in \mathbb{Z}^2} \rho_T(h)$  of the correlation function of  $Z_T$  is called the **integral range** of  $Z_T$ . This is a dimensionless quantity that satisfies  $0 \leq a_T \leq \infty$ .

**Property:**

If  $0 < a_T < \infty$ , and if  $\#V \gg a_T$ , then  $N_T(V)$  is approximately **Gaussianly distributed** with mean  $\#V \mu_T$  and variance  $\sigma_T^2 a_T \#V$

# Application to the choice of $V$

Put  $N_T(V) \approx \#V \mu_T + \sigma_T \sqrt{\#V a_T} Y$ , where  $Y$  is a standard Gaussian variable. Accordingly, we have

$$P\{N_T(V) \geq n\} \geq 1 - \alpha \iff P\left\{Y \geq \frac{n - \#V \mu_T}{\sigma_T \sqrt{\#V a_T}}\right\} \geq 1 - \alpha$$

Denoting by  $y_{1-\alpha}$  the quantile of order  $1 - \alpha$  of  $Y$ , the latter condition will be satisfied as soon as

$$\frac{n - \#V \mu_T}{\sigma_T \sqrt{\#V a_T}} \leq y_{1-\alpha},$$

which yields

$$\sqrt{\#V} \geq \frac{\sqrt{(1 - \mu_T) a_T y_{1-\alpha}^2} + \sqrt{(1 - \mu_T) a_T y_{1-\alpha}^2 + 4n}}{2\sqrt{\mu_T}}$$

The right handside member is a decreasing function of  $\mu_T$  and an increasing function of  $a_T$ .

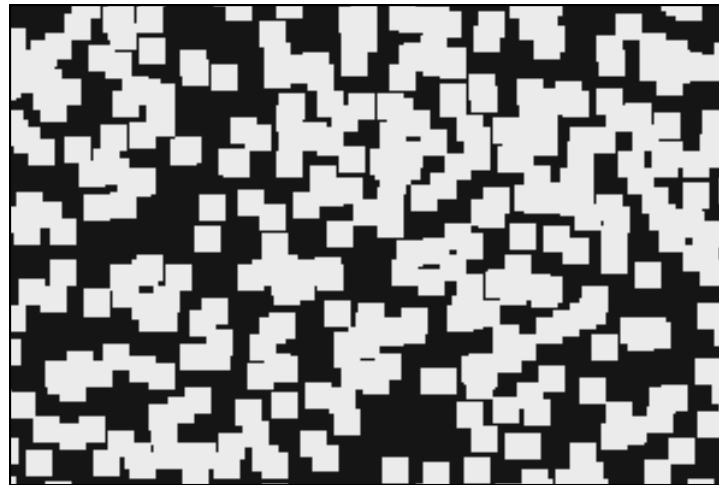
# Example: the discrete Boolean model

## Ingredients:

- Independent Poisson variables  $(N(u), u \in \mathbb{Z}^2)$  (mean value  $\theta$ );
- Independent copies  $(A_{u,n}, u \in \mathbb{Z}^2, n \leq N(u))$  of a random object  $A$ .

## Definition:

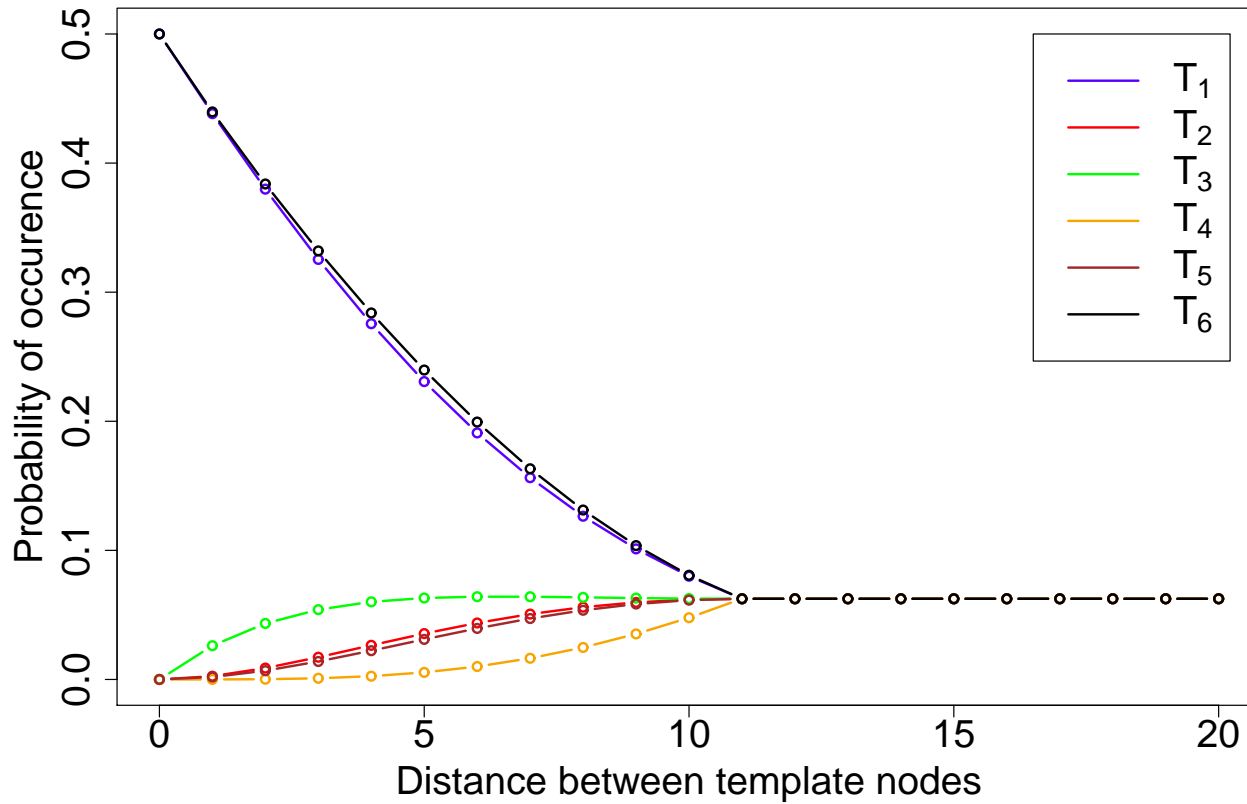
$$Z(x) = \max_{u \in \mathbb{Z}^2} 1_{x \in \tau_u A_u} \quad A_u = \bigcup_{n \leq N(u)} A_{u,n}$$



Boolean model of squares of side 11.  $\theta = 0.0057$  yields 50% zero proportion.

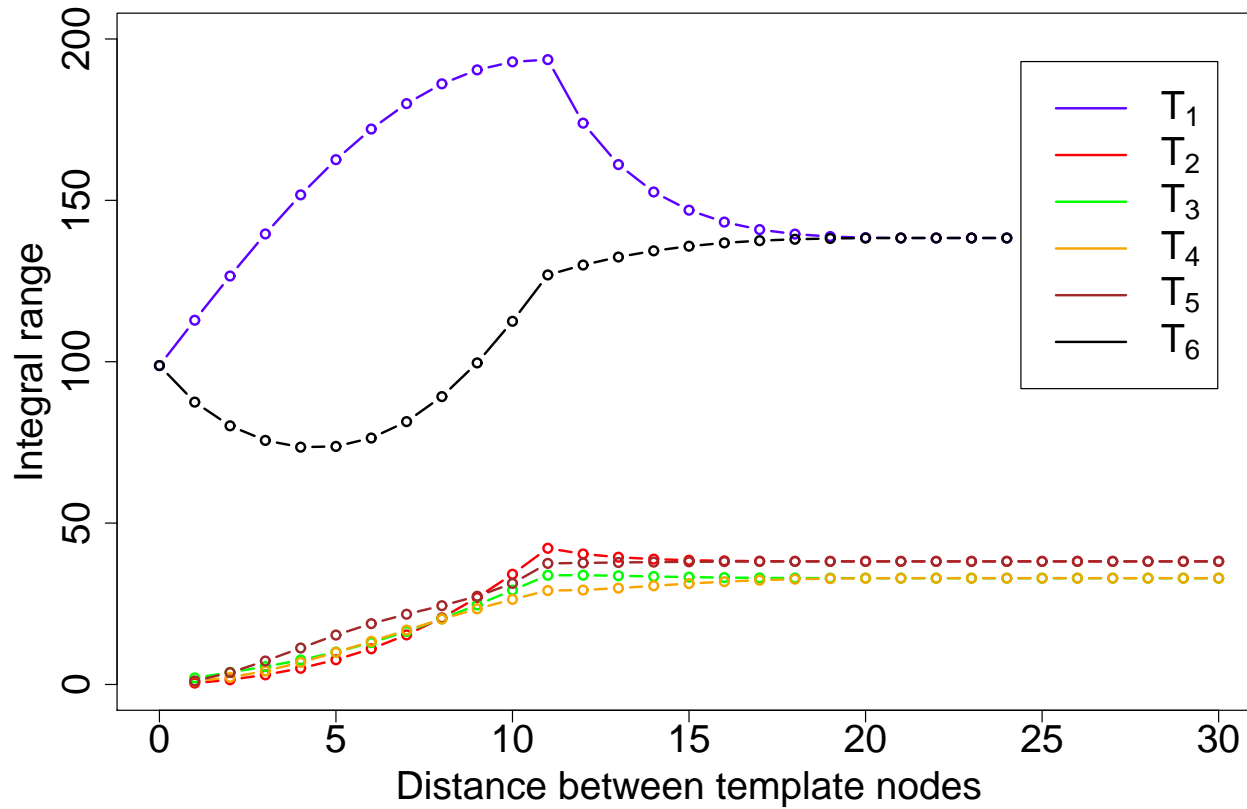
# Probability of matching

$$T_1 = \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix} \quad T_2 = \begin{pmatrix} 1 & 0 \\ 0 & 0 \end{pmatrix} \quad T_3 = \begin{pmatrix} 1 & 1 \\ 0 & 0 \end{pmatrix} \quad T_4 = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} \quad T_5 = \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} \quad T_6 = \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix}$$



# Integral range

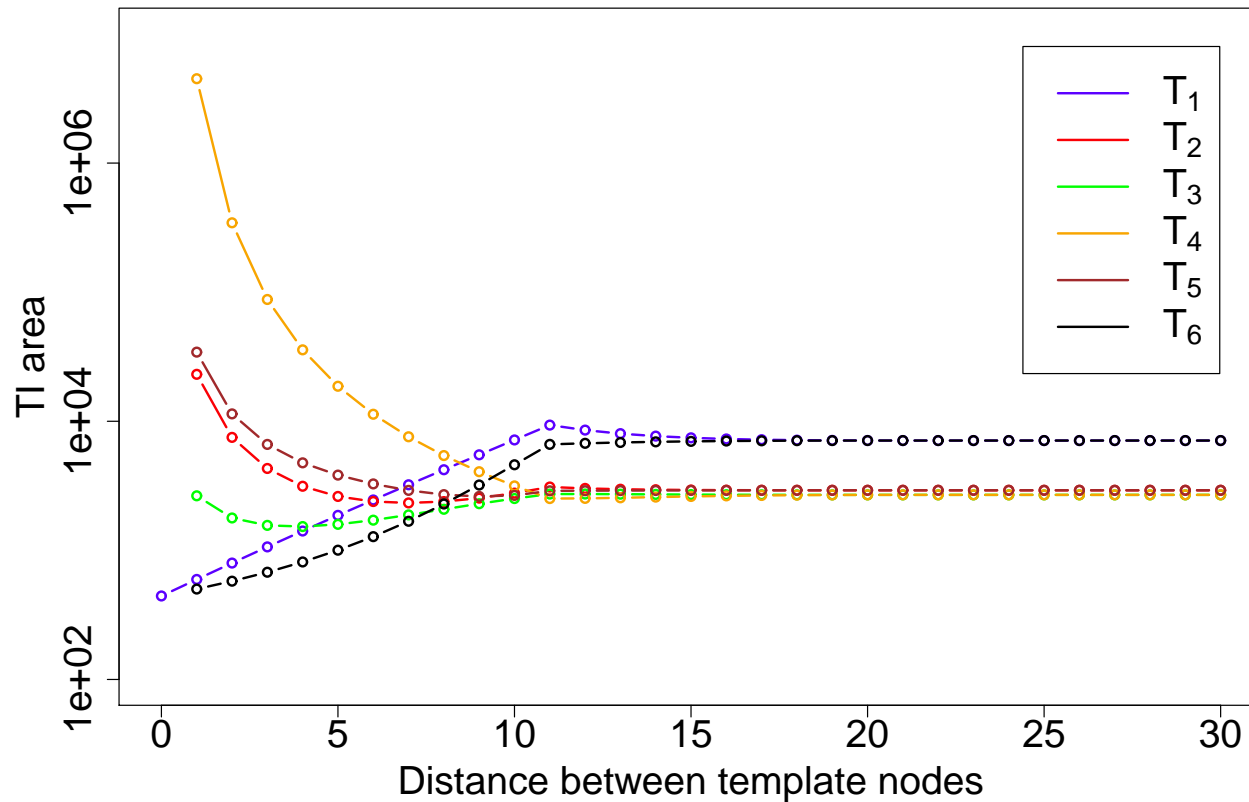
$$T_1 = \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix} \quad T_2 = \begin{pmatrix} 1 & 0 \\ 0 & 0 \end{pmatrix} \quad T_3 = \begin{pmatrix} 1 & 1 \\ 0 & 0 \end{pmatrix} \quad T_4 = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} \quad T_5 = \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} \quad T_6 = \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix}$$





# Required area for 50 matchings in 95% cases

$$T_1 = \begin{pmatrix} 0 & 0 \\ 0 & 0 \end{pmatrix} \quad
 T_2 = \begin{pmatrix} 1 & 0 \\ 0 & 0 \end{pmatrix} \quad
 T_3 = \begin{pmatrix} 1 & 1 \\ 0 & 0 \end{pmatrix} \quad
 T_4 = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix} \quad
 T_5 = \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix} \quad
 T_6 = \begin{pmatrix} 1 & 1 \\ 1 & 1 \end{pmatrix}$$



# A simple combinatorial remark

## Assumptions:

- The training image is a square of  $n^2$  pixels;
- The population of templates considered have the **same support** of  $k$  pixels.

## Counting:

- The total number of templates of the population is  $2^k$ .
- The training image contains at most  $n^2$  different templates of the population (independent of  $k!$ );

## Conclusion:

- The proportion of templates present in the training image is at most  $n^2/2^k$ .
- To give an order of magnitude,  $n = 10,000$  and  $k = 100$  (square  $10 \times 10$ ) yields an upper bound of  $8 \times 10^{-23}$  for the proportion, that is close to the reciprocal of the Avogadro number...